## Analysis of Algorithms The Mathematics of Analysis of Algorithms

Andres Mendez-Vazquez

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## Outline

#### Induction

- Basic Induction
- Structural Induction

#### Series

#### Properties

- Important Series
- Bounding the Series

#### Probability

- Intuitive Formulation
- Axioms
- Independence
- Unconditional and Conditional Probability
- Posterior (Conditional) Probability
- Random Variables
- Types of Random Variables
- Cumulative Distributive Function
- Properties of the PMF/PDF
- Expected Value and Variance
- Indicator Random Variable



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#### Principle of Mathematical Induction

Let P(n) be a property that is defined for integers n, and let a be a fixed integer.



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#### Suppose the following two statements are true

**1** P(a) is true.

For all integers  $k \ge a$ , if P(k) is true then P(k+1) is true.



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#### Principle of Mathematical Induction

Let P(n) be a property that is defined for integers n, and let a be a fixed integer.

#### Suppose the following two statements are true

• P(a) is true.

**2** For all integers  $k \ge a$ , if P(k) is true then P(k+1) is true.

# For all integers $n \ge a, P(n)$ is true.



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#### Then the statement

For all integers  $n \ge a$ , P(n) is true.

(1)

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## We have the following method for Mathematical Induction

#### Consider a statement of the form

For all integers  $n \ge a$ , P(n) is true.

#### To prove such a statement

Perform the following two steps

Step 1 (Basis step)

Show that P(a) is true - we normally use a=1



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### Step 2 (Inductive step)

Show that for all integers  $k \ge a$ , if P(k) is true, then P(k+1) is true.

#### Inductive hypothesis

Suppose that P(k) is true, where k is any particular but arbitrarily chosen integer with  $k\geq a.$ 

Then, you can prove that

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#### Proposition

For all integers  $n \ge 8$ ,  $n \diamondsuit$  can be obtained using  $3 \diamondsuit$  and  $5 \diamondsuit$  coins.

#### Show that P(8) is true

P(8) is true because 8¢ can be obtained using one coin 3¢ and another coin of 5¢.

Show that for all integers  $k \ge 8$ , if P(k) is true then P(k + 1) is also true

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### Case 1 - There is a $5 \diamondsuit$ among those making the change for $k \diamondsuit$

In this case, replace  $5 \ensuremath{\diamondsuit}$  by two  $3 \ensuremath{\diamondsuit}.$  Thus, we get the change for  $(k+1) \ensuremath{\,\diamondsuit}$ 

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## We can go further...

#### Recursively defined sets and structures

Assume S is a set. We use two steps to define the elements of S.

#### Basis Step

Specify an initial collection of elements.

#### Recursive Step

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#### Let S be a set that has been defined recursively

#### To prove that every object in ${\boldsymbol{S}}$ satisfies a certain property:

Show that each object in the BASE for S satisfies the property.
 Show that for each rule in the RECURSION, if the rule is applied to objects in S that satisfy the property, then the objects defined by the rule also satisfy the property.



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## Example: Binary trees recursive definition

Recall that the set  ${\cal B}$  of binary trees over an alphabet  ${\cal A}$  is defined as follows

- Basis:  $\langle \rangle \in B$
- **2** Recursive definition: If  $L, R \in B$  and  $x \in A$  then  $\langle L, x, R \rangle \in B$ .

#### Now define the function $f:B ightarrow\mathbb{N}$ defined as

# $$\begin{split} f\left(\left<\right>\right) = 0 \\ f\left(\left< L, x, R \right>\right) = \begin{cases} 1 & \text{if } L = R = \left<\right> \\ f\left(L\right) + f\left(R\right) & \text{otherwise} \end{cases} \end{split}$$

Theorem

Let T in B be a binary tree. Then f(T) yields the number of leaves of  $\mathfrak{I}$ 

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## Proof

#### By structural induction on T

Basis: The empty tree has no leaves, so  $f(\langle \rangle) = 0$  is correct.

#### Induction

Let L, R be trees in  $B, x \in A$ .

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Suppose that f(L) and f(R) denotes the number of leaves of L and R, respectively.



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If  $L = R = \langle \rangle$ , then  $\langle L, x, R \rangle = \langle \langle \rangle, x, \langle \rangle \rangle$  has one leaf, namely x, so  $f(\langle \langle \rangle, x, \langle \rangle \rangle) = 1$  is correct.

#### Case 2

If L and R are both not empty, then the number of leaves of the tree  $\langle L,x,R\rangle$  is equal to the number of leaves of L plus the number of leaves of R.

$$f\left(\langle L, x, R \rangle\right) = f\left(L\right) + f\left(R\right)$$

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# Introduction

# When an algorithm contains an iterative control construct such as a while or for loop

It is possible to express its running time as a series:

 $\sum_{j=1} j$ 

#### Thus, what is the objective of using these series

To be able to find bounds for the complexities of algorithms



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### Definition

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$$a_1 + a_2 + \dots + a_n = \sum_{k=1}^n a_k \tag{5}$$

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### In the case of infinite series

$$a_1 + a_2 + \dots = \sum_{k=1}^{\infty} a_k = \lim_{n \to \infty} \sum_{k=1}^n a_k$$
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# Linearity

# For any real number c and any finite sequences $a_1, a_2, ..., a_n$ and $b_1, b_2, ..., b_n$

$$\sum_{k=1}^{n} [ca_k + b_k] = c \sum_{k=1}^{n} a_k + \sum_{k=1}^{n} b_k$$

#### For More

Please take a look at page 1146 of Cormen's book.



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Imagine that each term in the sum has the following structure:

$$\frac{1}{k} - \frac{1}{k+1} =$$

What is the result of the following sum?  

$$\sum_{k=1}^{n} \frac{1}{k(k+1)}$$
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## Definition

For any sequence  $a_0, a_1, ..., a_n$ ,

$$\sum_{k=1}^{n} (a_k - a_{k-1}) = a_n - a_0 \tag{10}$$

#### Similarly

$$\sum_{k=0}^{n-1} \left( a_k - a_{k+1} \right) = a_0 - a_n \tag{11}$$

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# Arithmetic series

# Summing over the set $\{1, 2, 3, ..., n\}$

We can prove that

$$\sum_{i=1}^{n} i = \frac{n(n+1)}{2}$$

Proof

Basis: If n = 1 then  $\frac{1 \times 2}{2} = 1$ 



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$$= \frac{(n+1)(n+2)}{2}$$



# Series of Squares and Sums

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$$\sum_{k=0}^{n} k^2 = \frac{n(n+1)(2n+1)}{6}$$
(13)

Series of Cubes

 $\sum_{k=0}^{n} k^2 = \frac{n^2 \left(n+1\right)^2}{4} \tag{14}$ 

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# Geometric Series

## Definition

For a real  $x \neq 1$ , we have that

$$\sum_{k=0}^{n} x^{k} = 1 + x + x^{2} + \dots + x^{n}$$

It is called the geometric series

#### It is possible to prove that

$$\sum_{k=0}^{n} x^{k} = \frac{x^{n+1} - 1}{x - 1} \tag{16}$$

(15)

#### Proof

Now multiply both sides by of (Eq. 15) by x

$$x\left[\sum_{k=0}^{n} x^{k}\right] = x + x^{2} + x^{3} + \dots + x^{n+1}$$
(17)  
25/97

# Geometric Series

### Definition

For a real  $x \neq 1$ , we have that

$$\sum_{k=0}^{n} x^{k} = 1 + x + x^{2} + \dots + x^{n}$$
(15)

It is called the geometric series

### It is possible to prove that

$$\sum_{k=0}^{n} x^{k} = \frac{x^{n+1} - 1}{x - 1} \tag{16}$$

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(17)
<sub>25/97</sub>

# Subtract (Eq. 17) from (Eq. 15)

$$\sum_{k=0}^{n} x^{k} - x \left[ \sum_{k=0}^{n} x^{k} \right] = 1 - x^{n+1}$$
(18)

Finally

$$\sum_{k=0}^{n} x^{k} = \frac{1 - x^{n+1}}{1 - x} = \frac{x^{n+1} - 1}{x - 1}$$



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(19)



# Infinite Geometric Series

## When the summation is infinite and $\left|x\right| < 1$

$$\sum_{k=0}^{\infty} x^k = \frac{1}{1-x}$$

Proof

Given that



(20)

# Infinite Geometric Series

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Proof Given that  $\sum_{k=0}^{\infty} x^k = 1$ 



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(21)



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# For more on the series

Please take a look to

Cormen's book - Appendix A.



# Outline

#### Induction

Basic Induction

Structural Induction

#### Series

#### Properties

- Important Series
- Bounding the Series

#### Probabili

- Intuitive Formulation
- Axioms
- Independence
- Unconditional and Conditional Probability
- Posterior (Conditional) Probability
- Random Variables
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- Cumulative Distributive Function
- Properties of the PMF/PDF
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This is quite useful for Analysis of Algorithms

#### Important

The most basic way to evaluate a series is to use mathematical induction.

#### Example

Prove that



This is quite useful for Analysis of Algorithms

#### Important

The most basic way to evaluate a series is to use mathematical induction.

# Example Prove that $\sum_{k=0}^{n} 3^{k} \leq c 3^{n} \tag{22}$



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# Fast Bounding of Series

## A quick upper bound on the arithmetic series

$$\sum_{k=1}^{n} k \le \sum_{k=1}^{n} n = n^2$$
(23)

#### general, for a series $\sum_{k=1}^n a_k$

If  $a_{\max} = \max_{1 \leq k \leq n} a_k$  then

$$\sum_{k=1}^{n} a_k \le n \cdot a_{\max}$$

#### Another fast way of bounding finite series is

Suppose that  $rac{a_{k+1}}{a_k} \leq r$  for all  $k \geq 0$  where 0 < r < 1

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Suppose that  $\frac{a_{k+1}}{a_k} \leq r$  for all  $k \geq 0$  where 0 < r < 1.

# Thus, we have

$$a_k \leq a_0 r^k$$

#### Thus, we can use a infinite decreasing geometric series

(25)

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## Thus, we can use a infinite decreasing geometric series

$$\sum_{k=0}^{n} a_k \le \sum_{k=0}^{\infty} a_0 r^k$$

(25)

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## Thus, we can use a infinite decreasing geometric series

$$\sum_{k=0}^{n} a_k \le \sum_{k=0}^{\infty} a_0 r^k$$
$$= a_0 \sum_{k=0}^{\infty} r^k$$

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## Thus, we have

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## Thus, we can use a infinite decreasing geometric series

$$\sum_{k=0}^{n} a_k \leq \sum_{k=0}^{\infty} a_0 r^k$$
$$= a_0 \sum_{k=0}^{\infty} r^k$$
$$= a_0 \frac{1}{1-r}$$

# Approximation by integrals

# When a summation has the from $\sum_{k=m}^{n} f(k)$ , where f(k) is a monotonically increasing function

$$\int_{m-1}^{n} f(x) \, dx \le \sum_{k=m}^{n} f(k) \le \int_{m}^{n+1} f(x) \, dx \tag{26}$$

# For example

# Given $\ln(n+1) = \int_{1}^{n+1} \frac{1}{x} dx \le \sum_{k=1}^{n} \frac{1}{k}$ (27)

#### In addition





# For example

# Given

$$\ln\left(n+1\right) = \int_{1}^{n+1} \frac{1}{x} dx \le \sum_{k=1}^{n} \frac{1}{k}$$
(27)

# In addition

$$\sum_{k=1}^{n} \frac{1}{k} \le \int_{1}^{n} \frac{1}{x} dx = \ln n$$
(28)



# Outline

#### Induction

Basic Induction

Structural Induction

#### Series

#### Properties

- Important Series
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#### Probability

#### Intuitive Formulation

- Axioms
- Independence
- Unconditional and Conditional Probability
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- Random Variables
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# Gerolamo Cardano: Gambling out of Darkness

# Gambling

Gambling shows our interest in quantifying the ideas of probability for millennia, but exact mathematical descriptions arose much later.

#### Gerolamo Cardano (16th century)

While gambling he developed the following rule!!!

#### Equal conditions

"The most fundamental principle of all in gambling is simply equal conditions, e.g. of opponents, of bystanders, of money, of situation, of the dice box and of the dice itself. To the extent to which you depart from that equity, if it is in your opponent's favour, you are a fool, and if in your own, you are unjust."

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# Gerolamo Cardano's Definition

## Probability

"If therefore, someone should say, I want an ace, a deuce, or a trey, you know that there are 27 favourable throws, and since the circuit is 36, the rest of the throws in which these points will not turn up will be 9; the odds will therefore be 3 to 1."

#### Meaning

Probability as a ratio of favorable to all possible outcomes!!! As long all events are equiprobable...

#### Thus, we get

P(AII favourable throws) =

umber All favourable throw: Number of All throws

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## **Empiric Definition**

Intuitively, the probability of an event A could be defined as:

$$P(A) = \lim_{n \to \infty} \frac{N(A)}{n}$$

Where N(A) is the number that event a happens in n trials.

#### Example

Imagine you have three dices, then



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- The total number of outcomes is 6<sup>3</sup>
- If we have event  ${\cal A}=$  all numbers are equal,  $|{\cal A}|=6$



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## Example

Imagine you have three dices, then

- The total number of outcomes is 6<sup>3</sup>
- If we have event  ${\cal A}=$  all numbers are equal,  $|{\cal A}|=6$
- Then, we have that  $P(A) = \frac{6}{6^3} = \frac{1}{36}$

# Outline

#### Induction

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Structural Induction

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#### Properties

- Important Series
- Bounding the Series

#### 3 Probability

Intuitive Formulation

#### Axioms

- Independence
- Unconditional and Conditional Probability
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## Axioms

Given a sample space S of events, we have that

• If  $A_1, A_2, ..., A_n$  are mutually exclusive events (i.e.  $P(A_i \cap A_j) = 0$ ) then:

 $P(A_1 \cup A_2 \cup \dots \cup A_n) = \sum^n P(A_i)$ 



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# Set Operations

## We are using

Set Notation

Thus

What Operations?



# Set Operations

We are using

Set Notation

## Thus

What Operations?



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# Example

# Setup

Throw a biased coin twice



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#### We have the following event

At least one head!!! Can you tell me which events are part of it?

What about this one?

Tail on first toss.
### Setup

Throw a biased coin twice



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#### We have the following event

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### We have four main methods of counting

- Ordered samples of size r without replacement
- Unordered samples of size r without replacement
- Unordered samples of size r with replacement.



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### We have four main methods of counting

- **2** Ordered samples of size r without replacement
  - Unordered samples of size r without replacement
  - Unordered samples of size r with replacement



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### We have four main methods of counting

- O Ordered samples of size r without replacement
- **③** Unordered samples of size r without replacement

Unordered samples of size *r* with replacement



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### We have four main methods of counting

- **②** Ordered samples of size r without replacement
- O Unordered samples of size r without replacement
- O Unordered samples of size r with replacement

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### Ordered samples of size r with replacement

### Definition

The number of possible sequences  $(a_{i_1},...,a_{i_r})$  for n different numbers is

$$n \times n \times \dots \times n = n^r \tag{30}$$

#### Example

If you throw three dices you have 6 imes 6 imes 6=216



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If you throw three dices you have  $6\times6\times6=216$ 



### Ordered samples of size r without replacement

### Definition

The number of possible sequences  $(a_{i_1},...,a_{i_r})$  for n different numbers is

$$n \times n - 1 \times ... \times (n - (r - 1)) = \frac{n!}{(n - r)!}$$
 (31)

#### Example

The number of different numbers that can be formed if no digit can be repeated. For example, if you have 4 digits and you want numbers of size 3.



### Ordered samples of size r without replacement

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The number of possible sequences  $(a_{i_1}, ..., a_{i_r})$  for n different numbers is

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Unordered samples of size r without replacement

### Definition

Actually, we want the number of possible unordered sets.





Unordered samples of size r without replacement

#### Definition

Actually, we want the number of possible unordered sets.

#### However

We have  $\frac{n!}{(n-r)!}$  collections where we care about the order. Thus  $\frac{\frac{n!}{(n-r)!}}{r!} = \frac{n!}{r!(n-r)!} = \binom{n}{r}$ (32)



Unordered samples of size r with replacement

### Definition

We want to find an unordered set  $\{a_{i_1}, ..., a_{i_r}\}$  with replacement

### Use a digit trick for that

Look at the Board

# $\binom{n+r-1}{r}$



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Unordered samples of size r with replacement

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Look at the Board

#### Thus

$$\left(\begin{array}{c} n+r-1\\ r \end{array}\right)$$



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(33)

### How?

### Change encoding by adding more signs

Imagine all the strings of three numbers with  $\{1, 2, 3\}$ 

#### We have

### Change encoding by adding more signs

Imagine all the strings of three numbers with  $\{1,2,3\}$ 

### We have

Old String	New String
111	1+0,1+1,1+2=123
112	1+0,1+1,2+2=124
113	1+0,1+1,3+2=125
122	1+0,2+1,2+2=134
123	1+0,2+1,3+2=135
133	1+0,3+1,3+2=145
222	2+0,2+1,2+2=234
223	2+0,2+1,3+2=235
233	1+0,3+1,3+2=245
333	3+0,3+1,3+2=345

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### Independence

### Definition

Two events A and B are independent if and only if  $P(A,B)=P(A\cap B)=P(A)P(B)$ 

#### Do you have any example?

Any idea?



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### We have two dices

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### We have the following events

• 
$$B = \{$$
First dice 3, 4 or 5 $\}$ 

Look at the board!!! Independence between A,B,C



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- $A = \{$ First dice 1,2 or 3 $\}$
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- $C = \{ \text{The sum of two faces is 9} \}$

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#### So, we can do

Look at the board!!! Independence between A, B, C



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We can use it to derive the Binomial Distribution

# WHAT ?????



### We have this

- "Success" has a probability *p*.
  - "Failure" has a probability 1-p.



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- Toss a coin independently n times.
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# Thus, taking a sample $\omega$

### $\omega = 11 \cdots 10 \cdots 0$

k 1's followed by n-k 0's.

#### We have then

 $P(\omega) = P(A_1 \cap A_2 \cap \ldots \cap A_k \cap A_{k+1}^c \cap \ldots \cap A_n^c)$ =  $P(A_1) P(A_2) \cdots P(A_k) P(A_{k+1}^c) \cdots P(A_n^c)$ =  $p^k (1-p)^{n-k}$ 

#### Important

The number of such sample is the number of sets with k elements.... or...



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$$\left( \begin{array}{c} n \\ k \end{array} \right)$$

## Did you notice?

### We do not care where the 1's and 0's are

Thus all the probabilities are equal to  $p^k (1-p)^k$ 

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$$\sum_{k \text{ 1's}} p\left(\omega^k\right)$$

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## Proving this is a probability

### Sum of these probabilities is equal to 1

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#### The other is simple

$$0 \le \binom{n}{k} p \left(1-p\right)^{n-k} \le 1 \ \forall k$$

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## Outline

#### Induction

Basic Induction

Structural Induction

#### Series

#### Properties

- Important Series
- Bounding the Series

#### B Probability

- Intuitive Formulation
- Axioms
- Independence

#### Unconditional and Conditional Probability

- Posterior (Conditional) Probability
- Random Variables
- Types of Random Variables
- Cumulative Distributive Function
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### Unconditional

This is the probability of an event A prior to arrival of any evidence, it is denoted by P(A). For example:

P(Cavity)=0.1 means that "in the absence of any other information, there is a 10% chance that the patient is having a cavity".





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 P(Cavity/Toothache)=0.8 means that "there is an 80% chance that the patient is having a cavity given that he is having a toothache"



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### Posterior Probabilities

Relation between conditional and unconditional probabilities

• Conditional probabilities can be defined in terms of unconditional probabilities:

$$P(A|B) = \frac{P(A,B)}{P(B)}$$

which generalizes to the chain rule P(A, B) = P(B)P(A|B) = P(A)P(B|A).



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#### Law of Total Probabilities

• if  $B_1, B_2, ..., B_n$  is a partition of mutually exclusive events and A is an event, then  $P(A) = \sum_{i=1}^n P(A \cap B_i)$ . An special case  $P(A) = P(A, B) + P(A, \overline{B})$ .



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- In addition, this can be rewritten into  $P(A) = \sum_{i=1}^{n} P(A|B_i)P(B_i)$ .





#### Three cards are drawn from a deck

Find the probability of no obtaining a heart

#### We have

- 52 cards
- 39 of them not a heart

#### Define

 $A_i = \{ Card i is not a heart \} Then?$ 





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## Independence and Conditional

#### From here, we have that...

P(A|B) = P(A) and P(B|A) = P(B).

#### Conditional independence

A and B are conditionally independent given C if and only if

### P(A|B, C) = P(A|C)

Example: P(WetGrass|Season, Rain) = P(WetGrass|Rain).



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## One Version

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$



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$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

- P(A) is the **prior probability** or marginal probability of A. It is "prior" in the sense that it does not take into account any information about B.
- P(A|B) is the conditional probability of A, given B. It is also called the posterior probability because it is derived from or depends upon the specified value of B.
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- *P*(*B*) is the **prior or marginal probability** of B, and acts as a normalizing constant.

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### General Form of the Bayes Rule

#### Definition

If  $A_1, A_2, ..., A_n$  is a partition of mutually exclusive events and B any event, then:

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### Setup

Throw two unbiased dice independently.

•  $A = \{$ sum of the faces  $= 8\}$ •  $B = \{$ faces are equal $\}$ 

Then calculate  $P\left(B|A\right)$ 

Look at the board



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### We have the following

#### Two coins are available, one unbiased and the other two headed

#### Assume

That you have a probability of  $\frac{3}{4}$  to choose the unbiased

- $\mathcal{B}_{[]} = \{ \mathsf{Unbiased coin chosen} \}$
- Biased coin chosen
  - Find that if a head come up, find the probability that the two headed com was chosen



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### Random Variables

#### Definition

In many experiments, it is easier to deal with a summary variable than with the original probability structure.



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## Random Variables

#### Definition

In many experiments, it is easier to deal with a summary variable than with the original probability structure.

#### Example

In an opinion poll, we ask 50 people whether agree or disagree with a certain issue.

- Suppose we record a "1" for agree and "0" for disagree.
- The sample space for this experiment has  $2^{50}$  elements. Why?
- Suppose we are only interested in the number of people who agree
- Define the variable X = number of "1" 's recorded out of 50.
- Easier to deal with this sample space (has only 51 elements).


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# It is necessary to define a function random variable as follow

#### $X:S\to \mathbb{R}$

Graphically



# Thus...

It is necessary to define a function random variable as follow

 $X:S\to \mathbb{R}$ 



#### How?

What is the probability function of the random variable is being defined from the probability function of the original sample space?

Suppose the range of the random variable X =< x<sub>1</sub>, x<sub>2</sub>, ..., x<sub>m</sub> >
Then, we observe X = x<sub>i</sub> if and only if the outcome of the random experiment is an s<sub>j</sub> ∈ S s.t. X(s<sub>j</sub>) = x<sub>j</sub> or



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$$P(X = x_j) = P(s_j \in S | X(s_j) = x_j)$$



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#### Setup

Throw a coin 10 times, and let R be the number of heads.

#### Then

 $S={
m all}$  sequences of length 10 with components H and T

We have for

 $\omega = \mathsf{H}\mathsf{H}\mathsf{H}\mathsf{H}\mathsf{T}\mathsf{T}\mathsf{H}\mathsf{T}\mathsf{H} \Rightarrow R\left(\omega\right) = 6$ 



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#### Setup

Let  ${\it R}$  be the number of heads in two independent tosses of a coin.

• Probability of head is .6

What are the probabilities?  $\Omega = \{HH, HT, TH, TT\}$ 

Thus, we can calculate  $P\left(R=0
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 $\Omega = \{\mathsf{HH},\mathsf{HT},\mathsf{TH},\mathsf{TT}\}$ 

Thus, we can calculate

P(R = 0), P(R = 1), P(R = 2)



#### Setup

Let  ${\it R}$  be the number of heads in two independent tosses of a coin.

• Probability of head is .6

What are the probabilities?

 $\Omega = \{\mathsf{HH},\mathsf{HT},\mathsf{TH},\mathsf{TT}\}$ 

Thus, we can calculate

 $P\left(R=0\right),P\left(R=1\right),P\left(R=2\right)$ 



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# Outline

Basic Induction

Structural Induction

#### Properties

- Important Series
- Bounding the Series

- Probability
- Intuitive Formulation
- Axioms
- Independence
- Unconditional and Conditional Probability
- Posterior (Conditional) Probability
- Random Variables

#### Types of Random Variables

- Cumulative Distributive Function
- Properties of the PMF/PDF
- Expected Value and Variance
- Indicator Random Variable



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# Types of Random Variables

#### Discrete

A discrete random variable can assume only a countable number of values.

#### Continuous

A continuous random variable can assume a continuous range of values.



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Probability Mass Function (PMF) and Probability Density Function (PDF)

The pmf /pdf of a random variable X assigns a probability for each possible value of X.





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#### Properties of the pmf and pdf

• Some properties of the pmf:

 $p_{x} p(x) = 1$  and  $P(a < X < b) = \sum_{k=a}^{b} p(k)$ .

In a similar way for the pdf:

 $\quad \int_{-\infty}^{\infty} p(x) dx = 1$  and  $P(a < X < b) = \int_{a}^{b} p(t) dt$  .



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#### Cumulative Distributive Function

- Properties of the PMF/PDF
- Expected Value and Variance
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## Cumulative Distribution Function

• With every random variable, we associate a function called Cumulative Distribution Function (CDF) which is defined as follows:

$$F_X(x) = P(f(X) \le x)$$

• With properties:

- $\models F_X(x) \ge 0$
- $F_X(x)$  in a non-decreasing function of X.



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Exclude • If X is discrete, its CDF can be computed as follows:  $P_X(x) = P(f(X) \le x) = \sum_{k=1}^{N} P(X_k = p_k).$ 

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#### Example

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#### 

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# Example: Discrete Function





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#### Continuous Function

If X is continuous, its CDF can be computed as follows:

$$F(x) = \int_{-\infty}^{x} f(t) dt.$$

#### Remark

Based in the fundamental theorem of calculus, we have the following equality.

$$p(x) = rac{dF}{dx}(x)$$

#### Note

This particular p(x) is known as the Probability Mass Function (PMF) or Probability Distribution Function (PDF).

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# Example: Continuous Function

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# Setup • A number X is chosen at random between a and b Xhas a uniform distribution • $f_X(x) = \frac{1}{b-a}$ for $a \le x \le b$

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# Setup • A number X is chosen at random between a and b Xhas a uniform distribution • $f_X(x) = \frac{1}{b-a}$ for $a \le x \le b$ • $f_X(x) = 0$ for x < a and x > bWe have $F_X(x) = P\left\{X \le x\right\} = \int^x f_X(t) dt$ (34) $P\left\{a < X \le b\right\} = \int_{a}^{b} f_X\left(t\right) dt$ (35)CES

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# Graphically

### Example uniform distribution



### Outline

#### Induction

Basic Induction

Structural Induction

#### Series

#### Properties

- Important Series
- Bounding the Series

#### 3 Probability

- Probability
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#### Properties of the PMF/PDF

- Expected Value and Variance
- Indicator Random Variable



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Properties of the PMF/PDF

#### Conditional PMF/PDF

We have the conditional pdf:

$$p(y|x) = \frac{p(x,y)}{p(x)}.$$

From this, we have the general chain rule

$$p(x_1, x_2, ..., x_n) = p(x_1 | x_2, ..., x_n) p(x_2 | x_3, ..., x_n) ... p(x_n).$$

Independence

If X and Y are independent, then:

$$p(x, y) = p(x)p(y).$$

Properties of the  $\mathsf{PMF}/\mathsf{PDF}$ 

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### Properties of the PMF/PDF

#### Law of Total Probability

$$p(y) = \sum_{x} p(y|x)p(x).$$

### Outline

#### Induction

Basic Induction

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#### Expected Value and Variance

Indicator Random Variable



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### Something Notable

You have the random variables  $R_1, R_2$  representing how long is a call and how much you pay for an international call:

if  $0 \le R_1 \le 3$ (minute)  $R_2 = 10$ (cents) if  $3 < R_1 \le 6$ (minute)  $R_2 = 20$ (cents) if  $6 < R_1 \le 9$ (minute)  $R_2 = 30$ (cents)

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We have then the probabilities  $P \{R_2 = 10\} = 0.6, P \{R_2 = 20\} = 0.25, P \{R_2 = 10\} = 0$ 

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 $P\{R_2 = 10\} = 0.6, P\{R_2 = 20\} = 0.25, P\{R_2 = 10\} = 0.15.$ 

If we observe N calls and N is very large

We can say that we have  $N \times 0.6$  calls and  $10 \times N \times 0.6$  the cost of those calls.

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### Similarly

- $\{R_2 = 20\} \Longrightarrow 0.25N$  and total cost 5N.
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#### The average

$$\frac{10(0.6N) + 20(.25N) + 30(0.15N)}{N} = 10(0.6) + 20(.25) + 30(0.15)$$
$$= \sum_{y} yP \{R_2 = y\}$$

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#### Definition

• Discrete random variable X:  $E(X) = \sum_{x} xp(x)$ .

Continuous random variable  $Y: E(Y) = \int_{x} x$ 



#### Definition

- Discrete random variable X:  $E(X) = \sum_{x} xp(x)$ .
- Continuous random variable Y:  $E(Y) = \int_x xp(x) dx$ .

### Extension to a function g(x)• $E(g(X)) = \sum_{x} g(x)p(x)$ (Discrete case). • $E(g(X)) = \int_{-\infty}^{+\infty} g(x)p(x)dx$ (Continuous case)



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# Linear Property

#### Linearity property of the Expected Value

### E(af(X) + bg(Y)) = aE(f(X)) + bE(g(Y))

Example for a discrete distribution





#### Linearity property of the Expected Value

E(af(X) + bg(Y)) = aE(f(X)) + bE(g(Y))

(36)

#### Example for a discrete distribution

$$E[aX + b] = \sum_{x} [ax + b] p(x|\theta)$$



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$$= aE[X] + b$$



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Imagine the following

We have the following functions

**9** g(x) = 0, x < 0



### Example

### Imagine the following

### We have the following functions

**1** 
$$f(x) = e^{-x}, x \ge 0$$

#### g(x) = 0, x < 0

The expected Value



### Example

### Imagine the following

We have the following functions

• 
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2 
$$g(x) = 0, x < 0$$

The expected Value



### Example

#### Imagine the following

We have the following functions

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$$f(x) = e^{-x}, x \ge 0$$

2 
$$g(x) = 0, x < 0$$

#### Find

The expected Value



### Variance

### Definition

• 
$$Var(X) = E((X - \mu)^2)$$
 where  $\mu = E(X)$ 

#### Standard Deviation

#### • The standard deviation is simply $\sigma = \sqrt{Var(X)}$



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### Indicator Random Variable

#### Definition

The indicator of an event A is a random variable  $I_A$  defined as follows:

$$I_{A}(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases}$$
(37)


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#### Thus

- If  $I_A = 1$  if A occurs.
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Because we can count events!!!



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### Why is this useful?

Because we can count events!!!

(37)

### The Expected value of an indicator function

# $E[I_A] = 0 \times P(I_A = 0) + 1 \times P(I_A = 1) = P(I_A = 1) = P(A)$ (38)



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## Method of Indicators

### It is possible to see random variables as a sum of indicators functions

$$R = I_{A_1} + \dots + I_{A_n}$$
(39)

#### Then

$$E[R] = \sum_{j=1}^{n} E[I_{A_j}] = \sum_{j=1}^{n} P(A_j)$$
(40)

#### Hopefully

It is easier to calculate  $P\left(A_{j}
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## Example

### A single unbiased die is tossed independently $\boldsymbol{n}$ times

- Let  $R_1$  be the numbers of 1's.
- Let  $R_2$  be the numbers of 2's.

#### Find $E[R_1R_2]$

We can express each variable as

$$R_1 = I_{A_1} + \dots + I_{A_n}$$
  
 $R_2 = I_{B_1} + \dots + I_{B_n}$ 

$$E\left[R_1R_2
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### Thus

$$E[R_1 R_2] = \sum_{i=1}^{n} \sum_{j=1}^{n} E\left[I_{A_i} I_{B_j}\right]$$
(41)

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## Next

## Case 1 i eq j $I_{A_i}$ and $I_{B_j}$ are independent

$$E\left[I_{A_{i}}I_{B_{j}}\right] = E\left[I_{A_{i}}\right]E\left[I_{B_{j}}\right] = P(A_{i})P(B_{j}) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$$
(42)

#### Case 2 $i = j A_i$ and $B_i$ are disjoint

Thus,  $I_{A_i}I_{B_i}=I_{A_i\cap B_i}=0$ 

#### Thus

$$E\left[R_1R_2\right] = \frac{n\left(n-1\right)}{36}$$



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## Next

### Case 1 $i \neq j \ I_{A_i}$ and $I_{B_j}$ are independent

$$E\left[I_{A_{i}}I_{B_{j}}\right] = E\left[I_{A_{i}}\right]E\left[I_{B_{j}}\right] = P(A_{i})P(B_{j}) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$$
(42)

## Case 2 $i = j A_i$ and $B_i$ are disjoint

Thus,  $I_{A_i}I_{B_i} = I_{A_i \cap B_i} = 0$ 





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## Next

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$$E\left[I_{A_{i}}I_{B_{j}}\right] = E\left[I_{A_{i}}\right]E\left[I_{B_{j}}\right] = P(A_{i})P(B_{j}) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$$
(42)

## Case 2 $i = j A_i$ and $B_i$ are disjoint

Thus,  $I_{A_i}I_{B_i} = I_{A_i \cap B_i} = 0$ 

$$E[R_1 R_2] = \frac{n(n-1)}{36}$$
(43)

